

# 第十四讲 $R^2$ 的分解

2020.4.3

$$R_{I \cup J}^2 = R_I^2 + (1 - R_I^2) R_{J \cdot I}^2$$

命题3. 对于模型

$$\mathbf{y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon} = \mathbf{1}\beta_0 + Z\boldsymbol{\gamma} + \boldsymbol{\varepsilon}, \quad \boldsymbol{\varepsilon} \sim (0, \sigma^2)$$

记样本方差-协方差矩阵:  $\hat{\Sigma}_{\mathbf{xx}} = Z_{(c)}^\top Z_{(c)} / (n-1)$ ,  $\hat{\Sigma}_{\mathbf{xy}} = Z_{(c)}^\top \mathbf{y} / (n-1)$ ,

以及  $\hat{\Sigma}_{\mathbf{yy} \bullet \mathbf{x}} = \hat{\Sigma}_{\mathbf{yy}} - \hat{\Sigma}_{\mathbf{yx}} \hat{\Sigma}_{\mathbf{xx}}^{-1} \hat{\Sigma}_{\mathbf{xy}}$ , 则  $R^2 = \frac{\hat{\Sigma}_{\mathbf{yx}} \hat{\Sigma}_{\mathbf{xx}}^{-1} \hat{\Sigma}_{\mathbf{xy}}}{\hat{\Sigma}_{\mathbf{yy}}}$ .

$$X = (\mathbf{1}, X_I, X_J)$$

命题4. 对于模型  $\mathbf{y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon} = \mathbf{1}\beta_0 + X_I\boldsymbol{\beta}_I + X_J\boldsymbol{\beta}_J + \boldsymbol{\varepsilon}$ ,  $\boldsymbol{\varepsilon} \sim (0, \sigma^2)$

其中  $I \subset \{1, 2, \dots, p-1\}$ ,  $J = \{1, 2, \dots, p-1\} \setminus I$ . 则  $R^2 = R_{I \cup J}^2$  可分解为

$$R_{I \cup J}^2 = \frac{\hat{\Sigma}_{\mathbf{y}, I} \hat{\Sigma}_{I, I}^{-1} \hat{\Sigma}_{I, \mathbf{y}}}{\hat{\Sigma}_{\mathbf{yy}}} + \frac{\hat{\Sigma}_{\mathbf{y}, J \bullet I} \hat{\Sigma}_{JJ \bullet I}^{-1} \hat{\Sigma}_{J \bullet I, \mathbf{y}}}{\hat{\Sigma}_{\mathbf{yy}}} \triangleq R_I^2 + (1 - R_I^2) R_{J \bullet I}^2$$

其中  $R_{J \bullet I}^2 = \frac{\hat{\Sigma}_{\mathbf{y}, J \bullet I} \hat{\Sigma}_{JJ \bullet I}^{-1} \hat{\Sigma}_{J \bullet I, \mathbf{y}}}{\hat{\Sigma}_{\mathbf{yy} \bullet I}}$ , 称为复偏相关系数平方, 其中  $\hat{\Sigma}_{\mathbf{y}, I}$  为  $\mathbf{y}$  与  $\mathbf{x}_I$  的样本

协方差阵,  $\hat{\Sigma}_{I, I}$  为  $\mathbf{x}_I$  的样本协方差矩阵等等.

定义：

(1)  $R_{J \bullet I}^2 = \frac{\hat{\Sigma}_{yJ \bullet I} \hat{\Sigma}_{JJ \bullet I}^{-1} \hat{\Sigma}_{Jy \bullet I}}{\hat{\Sigma}_{yy \bullet I}}$  称为复偏相关系数平方, 代表了自变量  $X_I$  所不能解释

的  $y$  的方差中能被  $X_J$  所解释的比率(即控制变量  $X_I$  后的残差对变量  $X_J$  回归的复相关系数平方).

(2)  $(1 - R_I^2) R_{J \bullet I}^2 = R_{J \cup I}^2 - R_I^2 = \frac{\hat{\Sigma}_{yJ \bullet I} \hat{\Sigma}_{JJ \bullet I}^{-1} \hat{\Sigma}_{Jy \bullet I}}{\hat{\Sigma}_{yy}}$  称为半偏相关系数平方 (semipartial

correlation), 代表了当自变量从  $X_I$  增加到  $X_{I \cup J} = Z$  时, 复相关系数平方的增量.

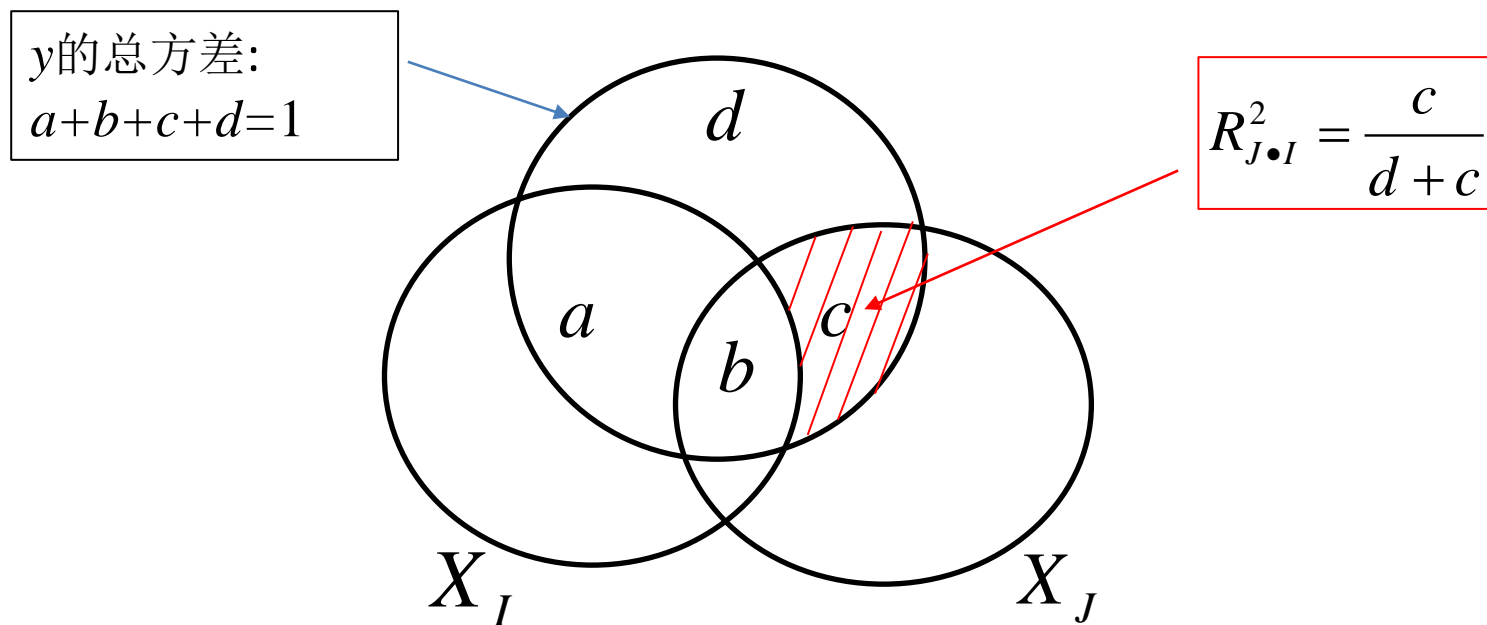
注：当  $J = \{j\}$  时,  $R_{j \bullet I}^2 = \frac{\hat{\Sigma}_{yj \bullet I} \hat{\Sigma}_{jj \bullet I}^{-1} \hat{\Sigma}_{jy \bullet I}}{\hat{\Sigma}_{yy \bullet I}} = (r_{yj \bullet I})^2$ ,  $r_{yj \bullet I}$  为  $y$  与  $x_j$  的偏相关系数.

$a + b + c = R^2 = R_{I \cup J}^2$ :  $X_I, X_J$ 所能解释的y的方差的比率;

$a + b = R_I^2$ :  $X_I$ 的贡献率(所能解释的y的方差的比率);

$c = R_{I \cup J}^2 - R_I^2$ :  $X_J$ 单独的贡献率;

$c / (d + c) = c / (1 - a - b) = R_{J \cdot I}^2$ : 排除 $X_I$ 的贡献( $a + b$ )之后,  $X_J$ 的贡献率。



$$R_{I \cup J}^2 = R_I^2 + (1 - R_I^2) R_{J \cdot I}^2 \Leftrightarrow a + b + c = (a + b) + (1 - a - b) \times \frac{c}{d + c}$$

证明：(1)不妨假设所有自变量  $Z$  已经中心化,

令  $X_J^\perp = X_J - P_{X_I} X_J$ , 我们有正交分解:

$$\hat{\mathbf{y}} - \mathbf{1}\bar{y} = P_{X_I} \mathbf{y} + P_{X_J^\perp} \mathbf{y} = P_{X_I} \mathbf{y} + X_J^\perp \hat{\boldsymbol{\beta}}_J,$$

$$\text{所以 } SS_{\square} = \|\hat{\mathbf{y}} - \mathbf{1}\bar{y}\|^2 = \|P_{X_I} \mathbf{y}\|^2 + \|P_{X_J^\perp} \mathbf{y}\|^2 = \|P_{X_I} \mathbf{y}\|^2 + \|X_J^\perp \hat{\boldsymbol{\beta}}_J\|^2,$$

注意

$$\|P_{X_I} \mathbf{y}\|^2 = \mathbf{y}^\top X_I (X_I^\top X_I)^{-1} X_I^\top \mathbf{y} = (n-1) \hat{\Sigma}_{y,I} \hat{\Sigma}_{I,I}^{-1} \hat{\Sigma}_{I,y},$$

由推论1,  $\hat{\boldsymbol{\beta}}_J = \hat{\Sigma}_{JJ\bullet-J}^{-1} \hat{\Sigma}_{Jy\bullet-J}$ , 并注意到  $X_J^{\perp\top} X_J^\perp = (n-1) \hat{\Sigma}_{JJ\bullet-J}$  所以

$$\|X_J^\perp \hat{\boldsymbol{\beta}}_J\|^2 = \hat{\boldsymbol{\beta}}_J^\top (X_J^{\perp\top} X_J^\perp) \hat{\boldsymbol{\beta}}_J = (n-1) \hat{\Sigma}_{yy\bullet-J} \hat{\Sigma}_{JJ\bullet-J}^{-1} \hat{\Sigma}_{Jy\bullet-J}$$

所以

$$\begin{aligned}
 R^2 &= R_{I \cup J}^2 = \frac{\|\hat{\mathbf{y}} - \mathbf{1}\bar{y}\|^2}{\|\mathbf{y} - \mathbf{1}\bar{y}\|^2} = \frac{\|P_{X_I}\mathbf{y}\|^2}{\|\mathbf{y} - \mathbf{1}\bar{y}\|^2} + \frac{\|X_J^\perp \hat{\boldsymbol{\beta}}_J\|^2}{\|\mathbf{y} - \mathbf{1}\bar{y}\|^2} \\
 &= \frac{\hat{\Sigma}_{y,I} \hat{\Sigma}_{I,I}^{-1} \hat{\Sigma}_{I,y}}{\hat{\Sigma}_{yy}} + \frac{\hat{\Sigma}_{yJ \bullet I} \hat{\Sigma}_{JJ \bullet I}^{-1} \hat{\Sigma}_{Jy \bullet I}}{\hat{\Sigma}_{yy}} \\
 &\triangleq R_I^2 + \Delta R^2
 \end{aligned}$$

$$\begin{aligned}
 (2) \quad \Delta R^2 &= \frac{\hat{\Sigma}_{yJ \bullet I} \hat{\Sigma}_{JJ \bullet I}^{-1} \hat{\Sigma}_{Jy \bullet I}}{\hat{\Sigma}_{yy}} = \frac{\hat{\Sigma}_{yJ \bullet I} \hat{\Sigma}_{JJ \bullet I}^{-1} \hat{\Sigma}_{Jy \bullet I}}{\hat{\Sigma}_{yy \bullet I}} \times \left( \frac{\hat{\Sigma}_{yy \bullet I}}{\hat{\Sigma}_{yy}} \right) \\
 &= R_{I \bullet J}^2 \times \left( \frac{\hat{\Sigma}_{yy} - \hat{\Sigma}_{y,I} \hat{\Sigma}_{I,I}^{-1} \hat{\Sigma}_{I,y}}{\hat{\Sigma}_{yy}} \right) \\
 &= R_{I \bullet J}^2 \times (1 - R_I^2)
 \end{aligned}$$

# 方差膨胀因子

自变量之间的相关性称为复共线性，  
复共线性对于LS估计的方差有何影响？

模型  $\mathbf{Y} = X\boldsymbol{\beta} + \boldsymbol{\varepsilon} = \mathbf{1}\beta_0 + Z\boldsymbol{\gamma} + \boldsymbol{\varepsilon}$ ,  $\boldsymbol{\varepsilon} \sim (0, \sigma^2)$ , 其中

$$\boldsymbol{\beta} = (\beta_0, \beta_1, \dots, \beta_{p-1})^\top,$$
$$X = (\mathbf{x}_{(0)} = \mathbf{1}, \mathbf{x}_{(1)}, \dots, \mathbf{x}_{(p-1)}),$$

由命题2,  $k \geq 1$ ,  $\mathbf{x}_{(k)}^\perp = \mathbf{x}_{(k)} - P_{X_{(-k)}} \mathbf{x}_{(k)}$ ,

$$\text{var}(\hat{\beta}_k | X) = \sigma^2 / \|\mathbf{x}_{(k)}^\perp\|^2 = \sigma^2 / [(n-1)\hat{\Sigma}_{\mathbf{x}_k \mathbf{x}_k \bullet \mathbf{x}_{-k}}].$$

$\|\mathbf{x}_{(k)}^\perp\|$  的大小依赖于  $\mathbf{x}_{(k)}$  与其它列的线性相关程度。

如果 $\mathbf{x}_k$ 与Z的其它列正交, 则  $\text{var}_0(\hat{\beta}_k | X) = \frac{\sigma^2}{\|\mathbf{x}_{(k)} - \mathbf{1}\bar{x}_{(k)}\|^2}$

如果 $\mathbf{x}_k$ 与其它列不正交, 则

$$\text{var}(\hat{\beta}_k | X) = \frac{\sigma^2}{\|\mathbf{x}_{(k)}^\perp\|^2} = \frac{\|\mathbf{x}_{(k)} - \mathbf{1}\bar{x}_{(k)}\|^2}{\|\mathbf{x}_{(k)}^\perp\|^2} \times \frac{\sigma^2}{\|\mathbf{x}_{(k)} - \mathbf{1}\bar{x}_k\|^2}$$

该方差比正交情形下的方差扩大的倍数为  $\|\mathbf{x}_{(k)} - \mathbf{1}\bar{x}_k\|^2 / \|\mathbf{x}_{(k)}^\perp\|^2$

定义: 对于带截距项的回归, 设计阵 $X = (\mathbf{x}_{(0)} = \mathbf{1}, \mathbf{x}_{(1)} = \dots, \mathbf{x}_{(p-1)})$ , 对 $1 \leq k \leq p-1$ , 定义方差膨胀因子(variance inflation factor)

$$\text{VIF}_k = \|\mathbf{x}_{(k)} - \mathbf{1}\bar{x}_k\|^2 / \|\mathbf{x}_{(k)}^\perp\|^2 = \hat{\Sigma}_{\mathbf{x}_k \mathbf{x}_k} / \hat{\Sigma}_{\mathbf{x}_k \mathbf{x}_k \bullet \mathbf{x}_{-k}}$$



$$\text{命题5. } \text{VIF}_k = \frac{1}{1 - R_k^2} \geq 1, R_k^2 = \frac{\|\hat{\mathbf{x}}_{(k)} - \mathbf{1}\bar{x}_k\|^2}{\|\mathbf{x}_{(k)} - \mathbf{1}\bar{x}_k\|^2} = \frac{\hat{\Sigma}_{k,-k} \hat{\Sigma}_{-k,-k}^{-1} \hat{\Sigma}_{-k,k}}{\hat{\Sigma}_{k,k}},$$

其中  $\hat{\mathbf{x}}_{(k)} = \mathbf{P}_{X_{(-k)}} \mathbf{x}_{(k)}$

证：因为  $\mathbf{x}_{(k)} = \hat{\mathbf{x}}_{(k)} \oplus \mathbf{x}_{(k)}^\perp$ ，且  $\mathbf{x}_{(k)}^\perp \perp \hat{\mathbf{x}}_{(k)}$ ， $\mathbf{x}_{(k)}^\perp \perp \mathbf{1}$ ，所以

$$(\mathbf{x}_{(k)} - \mathbf{1}\bar{x}_k) = (\hat{\mathbf{x}}_{(k)} - \mathbf{1}\bar{x}_k) \oplus \mathbf{x}_{(k)}^\perp$$

$$\text{所以 } \frac{1}{\text{VIF}_k} = \frac{\|\mathbf{x}_{(k)}^\perp\|^2}{\|\mathbf{x}_{(k)} - \mathbf{1}\bar{x}_k\|^2} = \frac{\|\mathbf{x}_{(k)} - \mathbf{1}\bar{x}_k\|^2 - \|\hat{\mathbf{x}}_{(k)} - \mathbf{1}\bar{x}_k\|^2}{\|\mathbf{x}_{(k)} - \mathbf{1}\bar{x}_k\|^2} = 1 - R_k^2$$

$$0 \leq R_k^2 \leq 1 \Rightarrow \text{VIF}_k \geq 1.$$

注：  $R_k^2$  为  $\mathbf{x}_{(k)}$  对其它自变量回归的复相关系数平方，度量了  $\mathbf{x}_{(k)}$  与其它自变量的相关性，故 VIF 与复共线性有关。

# 例子：两变量回归

我们以两个自变量的模型为例，考察LS估计以及 $R^2$ .

例1. 假设数据 $(y_i, x_{i1}, x_{i2}), i = 1, 2, \dots, n$ , 满足模型：

$$y_i = a + bx_{i1} + cx_{i2} + \varepsilon_i, \quad \varepsilon_i \text{ iid } \sim (0, \sigma^2), \quad \varepsilon_i \text{ 与 } x_{i1}, x_{i2} \text{ 独立。}$$

记号：

$$s_{12} = \sum (x_{i1} - \bar{x}_1)(x_{i2} - \bar{x}_2),$$

$$s_{11} = \sum (x_{i1} - \bar{x}_1)^2$$

$$s_{y1} = \sum (y_i - \bar{y})(x_{i1} - \bar{x}_1) \text{ 等等}$$

对于两变量回归，我们有如下结果：

$$(1) \hat{b} = \frac{s_{y1} - s_{y2}s_{12}/s_{22}}{s_{11} - s_{12}s_{21}/s_{22}} = \sqrt{\frac{s_{yy}}{s_{11}}} \left( \frac{r_{y1} - r_{y2}r_{12}}{1 - r_{12}^2} \right),$$

$$\text{var}(\hat{b}) = \frac{\sigma^2}{s_{11} - s_{12}^2/s_{22}} = \frac{\sigma^2}{s_{11}} \times \frac{1}{1 - r_{12}^2}, VIF = \frac{1}{1 - r_{12}^2}$$

$$(2) \Delta R^2 = R_{12}^2 - R_1^2 = \frac{(r_{y2} - r_{y1}r_{12})^2}{1 - r_{12}^2}, \quad R_{2 \bullet 1}^2 = \frac{(r_{y2} - r_{y1}r_{12})^2}{(1 - r_{y1}^2)(1 - r_{12}^2)} = (r_{y2 \bullet 1})^2,$$

其中  $r_{y2 \bullet 1} = \frac{r_{y2} - r_{y1}r_{12}}{\sqrt{(1 - r_{y1}^2)(1 - r_{12}^2)}}$  为  $y$  和  $x_2$  的偏相关系数.

$$(3) R^2 = R_1^2 + \Delta R^2 = \frac{r_{y1}^2 + r_{y2}^2 - 2r_{y1}r_{y2}r_{12}}{1 - r_{12}^2}.$$

(1) 由推论1,

$$\hat{b} = \frac{s_{y1\bullet 2}}{s_{11\bullet 2}} = \frac{s_{y1} - s_{y2}s_{12}/s_{22}}{s_{11} - s_{12}s_{21}/s_{22}} = \sqrt{\frac{s_{yy}}{s_{11}}} \left( \frac{r_{y1} - r_{y2}r_{12}}{1 - r_{12}^2} \right)$$

$$\text{var}(\hat{b}) = \frac{\sigma^2}{s_{11\bullet 2}} = \frac{\sigma^2}{s_{11} - s_{12}^2/s_{22}} = \frac{\sigma^2}{s_{11}} \times \frac{1}{1 - r_{12}^2}, \quad VIF = \frac{1}{1 - r_{12}^2}$$

$$R_{I \cup J}^2 = \frac{\hat{\Sigma}_{y,I} \hat{\Sigma}_{I,I}^{-1} \hat{\Sigma}_{I,y}}{\hat{\Sigma}_{yy}} + \frac{\hat{\Sigma}_{yJ\bullet I} \hat{\Sigma}_{JJ\bullet I}^{-1} \hat{\Sigma}_{Jy\bullet I}}{\hat{\Sigma}_{yy}} \triangleq R_I^2 + \Delta R^2 = R_I^2 + (1 - R_I^2) R_{J\bullet I}^2,$$

(2) 由命题4,  $R_1^2 = \frac{\hat{\Sigma}_{y1} \hat{\Sigma}_{11}^{-1} \hat{\Sigma}_{1y}}{\hat{\Sigma}_{yy}} = r_{y1}^2$

$$\begin{aligned}
\Delta R^2 &= R_{12}^2 - R_1^2 = \frac{\hat{\Sigma}_{y2\bullet 1} \hat{\Sigma}_{22\bullet 1}^{-1} \hat{\Sigma}_{2y\bullet 1}}{\hat{\Sigma}_{yy}} = \frac{(\hat{\Sigma}_{y2} - \hat{\Sigma}_{y1} \hat{\Sigma}_{11}^{-1} \hat{\Sigma}_{12})^2 (\hat{\Sigma}_{22} - \hat{\Sigma}_{21} \hat{\Sigma}_{11}^{-1} \hat{\Sigma}_{12})^{-1}}{\hat{\Sigma}_{yy}} \\
&= \frac{(r_{y2} - r_{y1} r_{12})^2}{1 - r_{12}^2} = \frac{(r_{y2} - r_{y1} r_{12})^2}{(1 - r_{y1}^2)(1 - r_{12}^2)} \times (1 - r_{y1}^2) \\
&= r_{y2\bullet 1}^2 (1 - r_{y1}^2)
\end{aligned}$$

偏相关系数(3个变量情形)

$$r_{12\bullet 3} = \frac{r_{12} - r_{13} r_{23}}{\sqrt{1 - r_{13}^2} \sqrt{1 - r_{23}^2}}$$

$$(3) \text{ 所以 } R^2 = R_1^2 + \Delta R^2 = r_{y1}^2 + \frac{(r_{y2} - r_{y1} r_{12})^2}{1 - r_{12}^2} = \frac{r_{y1}^2 + r_{y2}^2 - 2r_{y1} r_{y2} r_{12}}{1 - r_{12}^2}$$

总结如下:  $R_1^2 = r_{y1}^2$

$$\text{偏相关系数: } R_{2 \bullet 1} = r_{y2 \bullet 1} = \frac{r_{y2} - r_{y1}r_{12}}{\sqrt{1 - r_{12}^2} \sqrt{1 - r_{y1}^2}},$$

$$\text{半偏相关系数: } \Delta R^2 = \frac{r_{y2} - r_{y1}r_{12}}{\sqrt{1 - r_{12}^2}} = R_{2 \bullet 1} \sqrt{1 - r_{y1}^2},$$

$$R^2 = R_1^2 + \Delta R^2 = R_1^2 + R_{2 \bullet 1}(1 - R_1^2) = \frac{r_{y1}^2 + r_{y2}^2 - 2r_{y1}r_{y2}r_{12}}{1 - r_{12}^2}.$$

$$\text{对称地我们也有 } R_{12}^2 - R_2^2 = \frac{(r_{y1} - r_{y2}r_{12})^2}{1 - r_{12}^2} = r_{y1 \bullet 2}^2(1 - r_{y2}^2), \quad R_{1 \bullet 2}^2 = (r_{y1 \bullet 2})^2$$

$$R^2 = R_{12}^2 = r_{y2}^2 + \frac{(r_{y1} - r_{y2}r_{12})^2}{1 - r_{12}^2} = \frac{r_{y1}^2 + r_{y2}^2 - 2r_{y1}r_{y2}r_{12}}{1 - r_{12}^2}$$